

Answers to exercises

Chapter 1

Dioecious trees

The question is whether the number of flowers produced depends upon the sex of the tree. The variable *SEX* is categorical with two levels. The commands to execute the ANOVA are as before. We have assumed that the dataset *dioecioustrees* has been downloaded into the *gandh* library.

SAS COMMANDS FOR BOX 15.1 Analysis for dioecious trees	
Commands	<pre>proc glm data=gandh.dioecioustrees; class SEX; model FLOWERS = SEX; run;</pre>
Menu route	Statistics > Anova > Factorial Anova... FLOWERS → Dependent SEX → Independent

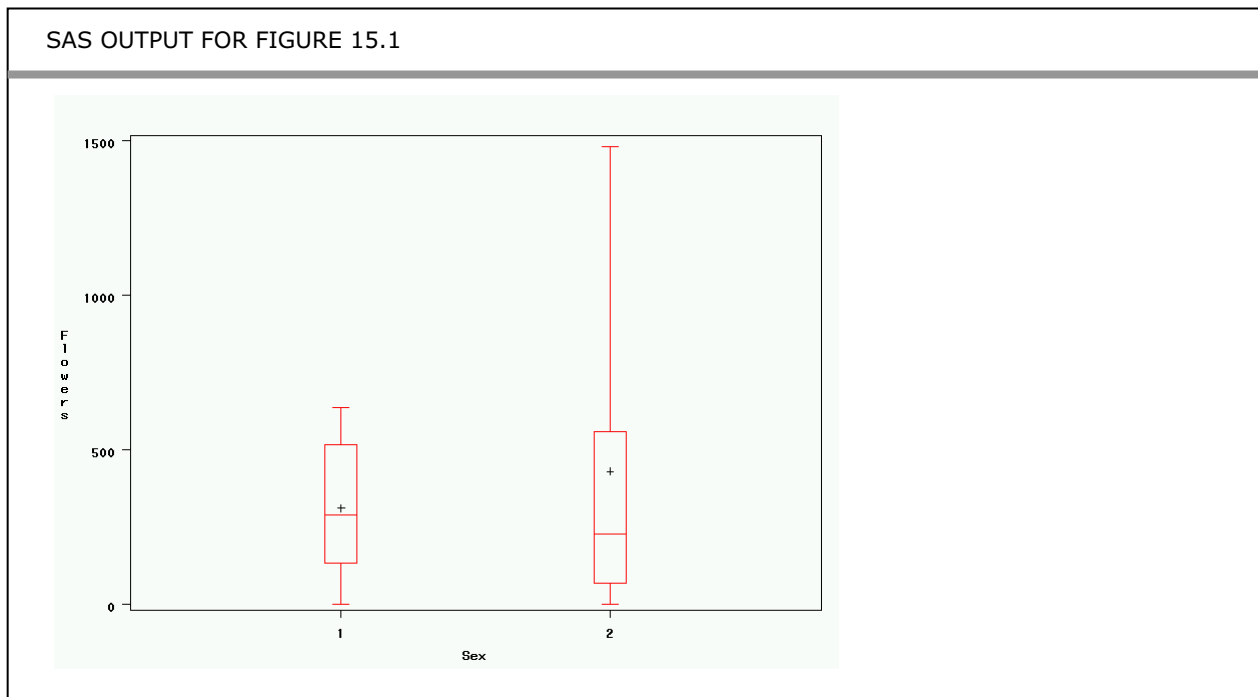
This would produce the following output:

SAS OUTPUT FOR BOX 15.1 Analysis for dioecious trees					
The GLM Procedure					
Dependent Variable: FLOWERS					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	171841.333	171841.333	1.18	0.2837
Error	48	7017255.167	146192.816		
Corrected Total	49	7189096.500			
	R-Square	Coef Var	Root MSE	FLOWERS Mean	
	0.023903	99.70057	382.3517	383.5000	
Source	DF	Type I SS	Mean Square	F Value	Pr > F
SEX	1	171841.3333	171841.3333	1.18	0.2837
Source	DF	Type III SS	Mean Square	F Value	Pr > F
SEX	1	171841.3333	171841.3333	1.18	0.2837

There are a number of ways in which these results could be illustrated graphically, one of which is a boxplot. Notice that in the command line route we need to sort the dataset by SEX in the first place, as SAS requires this for the grouping variable in the plot. The menu method handles this automatically.

SAS COMMANDS FOR FIGURE 15.1 Boxplot for dioecious trees	
Commands	<pre>proc sort; by SEX; proc boxplot; plot FLOWERS*SEX; run;</pre>
Menu route	Graphs > Box Plot FLOWERS → Analysis SEX → Class

This would produce the following graph:



Chapter 2

Dioecious trees

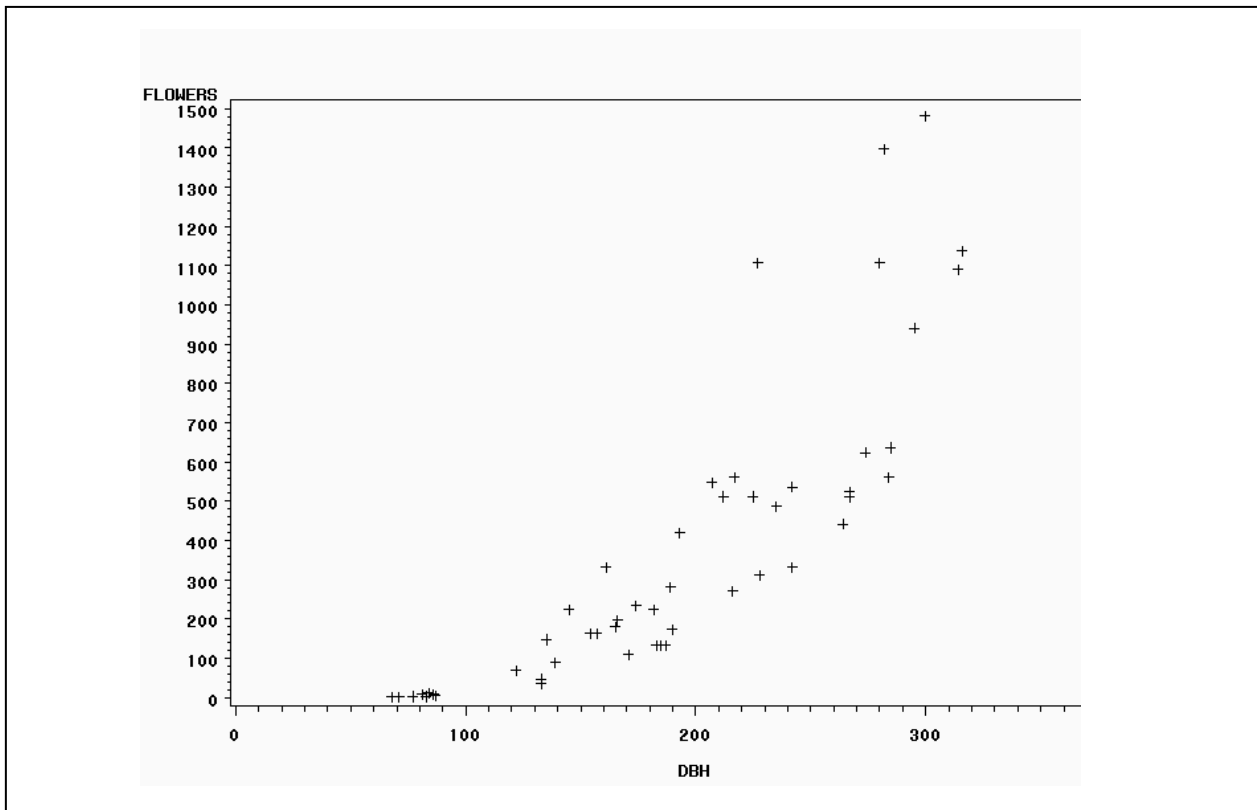
The variables FLOWERS and DBH are both continuous, so the best illustration of the relationship between them would be a scatter plot.

SAS COMMANDS FOR FIGURE 15.2 **Graph of FLOWERS versus DBH**

Commands `proc gplot data=gandh.Chapter2;`
`plot FLOWERS * DBH;`
`run;`

Menu route `Graphs > Scatter Plot > Two-Dimensional...`
 FLOWERS → Y Axis
 DBH → X Axis.

This would produce the following graph:



The regression analysis would be executed using the following commands:

SAS COMMANDS FOR BOX 15.2 **Analysis for dioecious trees**

```
Commands  proc reg data=gandh.Chapter2;
           model FLOWERS = DBH;
           run;
```

```
Menu route  Statistics > Regression > Linear...
           FLOWERS → Dependent
           DBH → Explanatory
```

which would give the following output:

SAS OUTPUT FOR BOX 15.2 Analysis for dioecious trees						
The REG Procedure						
Model: MODEL1						
Dependent Variable: FLOWERS						
Analysis of Variance						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	1	5060723	5060723	114.13	<.0001	
Error	48	2128374	44341			
Corrected Total	49	7189097				
	Root MSE	210.57330	R-Square	0.7039		
	Dependent Mean	383.50000	Adj R-Sq	0.6978		
	Coeff Var	54.90829				
Parameter Estimates						
Variable	Label	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	Intercept	1	-481.16042	86.24085	-5.58	<.0001
DBH	DBH	1	4.51284	0.42242	10.68	<.0001

Chapter 3

How variability in the population will influence our analysis

To create data sets for this exercise, just follow the procedure outlined in SAS Box 3.3a 'Creating your own datasets', but alter the standard deviation specified when you draw the random numbers from the Normal distribution (the 'sigma=' step in the second group of commands in SAS Box 3.3). We reproduce the commands in the following box, because here we have named the output file 'results'. One advantage of using the command line method is that you can repeatedly execute the whole lot by pressing the 'Run' button, just noting down the grand mean produced in the output box between runs. Then you can edit the number you want to change and repeat the whole exercise.

```

data withdums;
  set gandh.Chapter3sim;
  dum1=(FERTIL=1);
  dum2=(FERTIL=2);
  run;
data withparameters;
  set withdums;
  k3=13.5;
  k1=1.2;
  k2=3.4;
  sigma=8;
  run;
data results;
  set withparameters;
  noise=normal(0);
  Y=k3+k1*dum1+k2*dum2+sigma*noise;
  run;
proc glm data=work.results;
  class FERTIL;
  model Y=FERTIL / solution;
  run;

```

The final step may also be carried out by menu route as shown below:

Menu route Statistics > Anova > Linear Models...

Y → Dependent

FERTIL → Class

Statistics

Parameter estimates

Each analysis will produce one estimate of the grand mean, which may then be entered into a new column in a new worksheet. To create a new worksheet, make sure you are in the Analyst environment. Use File > New. The column headings may then be named SIGMA2, SIGMA4, SIGMA8 etc. The values obtained from each analysis may then be entered into this new worksheet. In the example below, this worksheet has been saved in the work library and called 'SIGS'. (An alternative way of creating this new dataset is explained for the menu route in the box below). When ten estimates have been obtained, a histogram may be drawn for the column SIGMA8 (for example) using the commands:

SAS COMMANDS FOR FIGURE 15.3

```

Commands  proc capability data=work.SIGS;
           var SIGMA8;
           histogram;
           run;

```

Menu route Edit > Insert Columns > Numeric

Data > Column Properties

SIGMA8 → Name

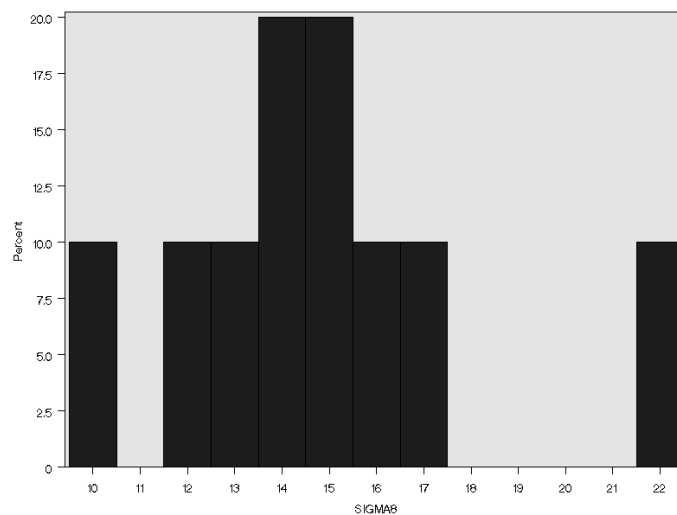
Now type the observed values of the grand mean into the column

Graphs > Histogram...

SIGMA8 → Analysis

This would produce a graph similar to the one below:

SAS OUTPUT FOR FIGURE 15.3



Comparison of the four graphs with the error standard deviation set at 2, 4, 8 and 16 will illustrate how the error variance will influence the reliability of the grand mean estimate (see main text for an example).

Chapter 4

Investigating Obesity

First you need to conduct two separate analyses to explain FOREARM using HT or WT.

SAS COMMANDS FOR BOX 15.4 (A) First analysis of FOREARM

```
Commands  proc glm data=gandh.Chapter4;
           model FOREARM = HT;
           run;
```

Menu route Statistics > Anova > Linear Models...
 FOREARM → Dependent
 HT → Quantitative

SAS OUTPUT FOR BOX 15.4 (A) First analysis of FOREARM

The GLM Procedure

Dependent Variable: FOREARM

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	0.9439390	0.9439390	0.18	0.6778
Error	37	199.0940026	5.3809190		
Corrected Total	38	200.0379416			

R-Square	Coeff Var	Root MSE	FOREARM Mean
0.004719	45.74012	2.319681	5.071436

Source	DF	Type I SS	Mean Square	F Value	Pr > F
HT	1	0.94393898	0.94393898	0.18	0.6778

Source	DF	Type III SS	Mean Square	F Value	Pr > F
HT	1	0.94393898	0.94393898	0.18	0.6778

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	0.9587990575	9.82623490	0.10	0.9228
HT	0.0260479467	0.06219132	0.42	0.6778

SAS COMMANDS FOR BOX 15.4 (B) **Second analysis of FOREARM**

```
Commands  proc glm data=gandh.Chapter4;
           model FOREARM = WT;
           run;
```

Menu route Statistics > Anova > Linear Models...

FOREARM → Dependent

WT → Quantitative

SAS OUTPUT FOR BOX 15.4 (B) **Second analysis of FOREARM**

The GLM Procedure

Dependent Variable: FOREARM

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	59.1372267	59.1372267	15.53	0.0003
Error	37	140.9007148	3.8081274		
Corrected Total	38	200.0379416			

	R-Square	Coeff Var	Root MSE	FOREARM Mean
	0.295630	38.47909	1.951442	5.071436

Source	DF	Type I SS	Mean Square	F Value	Pr > F
WT	1	59.13722674	59.13722674	15.53	0.0003

Source	DF	Type III SS	Mean Square	F Value	Pr > F
WT	1	59.13722674	59.13722674	15.53	0.0003

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	-4.594757196	2.47272868	-1.86	0.0711
WT	0.152978988	0.03882014	3.94	0.0003

Then the joint analysis using both explanatory variables

SAS COMMANDS FOR BOX 15.5 **Analysis of FOREARM using both explanatory variables**

```
Commands  proc glm data=gandh.Chapter4;
           model FOREARM = HT WT;
           run;
```

Menu route Statistics > Anova > Linear Models...

FOREARM → Dependent

HT WT → Quantitative

Statistics

Type 1

SAS OUTPUT FOR BOX 15.5 **Analysis of FOREARM using both explanatory variables**

The GLM Procedure					
Dependent Variable: FOREARM					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	83.9143634	41.9571817	13.01	<.0001
Error	36	116.1235782	3.2256550		
Corrected Total	38	200.0379416			
	R-Square	Coeff Var	Root MSE	FOREARM Mean	
	0.419492	35.41425	1.796011	5.071436	
Source	DF	Type I SS	Mean Square	F Value	Pr > F
HT	1	0.94393898	0.94393898	0.29	0.5919
WT	1	82.97042439	82.97042439	25.72	<.0001
Source	DF	Type III SS	Mean Square	F Value	Pr > F
HT	1	24.77713663	24.77713663	7.68	0.0088
WT	1	82.97042439	82.97042439	25.72	<.0001

Chapter 5

The dorsal crest of the male smooth newt

In the first analysis you are asked to investigate the relationship between LCREST and the body size of the newt. The response variable you are given is logged, therefore the body size variable should also be logged. The commands and menus are given below:

SAS COMMANDS FOR BOX 15.6 Dorsal crest analysis	
Commands	<pre>proc glm data=gandh.Chapter5; model LCREST = LSVL / solution; run;</pre>
Menu route	Statistics > Anova > Linear Models... LCREST → Dependent LSVL → Quantitative Statistics <input checked="" type="checkbox"/> Parameter estimates

This would result in the following output:

SAS OUTPUT FOR BOX 15.6 Dorsal crest analysis					
The GLM Procedure					
Dependent Variable: LCREST					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	2.38939159	2.38939159	45.81	<.0001
Error	85	4.43372434	0.05216146		
Corrected Total	86	6.82311593			
	R-Square	Coeff Var	Root MSE	LCREST Mean	
	0.350191	29.30913	0.228389	0.779241	
Source	DF	Type I SS	Mean Square	F Value	Pr > F
LSVL	1	2.38939159	2.38939159	45.81	<.0001
Source	DF	Type III SS	Mean Square	F Value	Pr > F
LSVL	1	2.38939159	2.38939159	45.81	<.0001
Parameter	Estimate	Standard Error	t Value	Pr > t	
Intercept	-9.380587570	1.50132730	-6.25	<.0001	
LSVL	5.086989483	0.75160915	6.77	<.0001	

In the second analysis, POND is also included.

SAS COMMANDS FOR BOX 15.7 Further dorsal crest analysis	
Commands	<pre>proc glm data=gandh.Chapter5; class POND; model LCREST = POND LSVL / solution; run;</pre>
Menu route	Statistics > Anova > Linear Models... LCREST → Dependent POND → Class LSVL → Quantitative <div style="border: 1px solid black; display: inline-block; padding: 2px;">Statistics</div> <input checked="" type="checkbox"/> Parameter estimates

SAS OUTPUT FOR BOX 15.7 Further dorsal crest analysis

The GLM Procedure

Dependent Variable: LCREST

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	10	2.63001908	0.26300191	4.77	<.0001
Error	76	4.19309685	0.05517233		
Corrected Total	86	6.82311593			

R-Square	Coeff Var	Root MSE	LCREST Mean
0.385457	30.14315	0.234888	0.779241

Source	DF	Type I SS	Mean Square	F Value	Pr > F
POND	9	0.32519109	0.03613234	0.65	0.7466
LSVL	1	2.30482798	2.30482798	41.78	<.0001

Source	DF	Type III SS	Mean Square	F Value	Pr > F
POND	9	0.24062749	0.02673639	0.48	0.8807
LSVL	1	2.30482798	2.30482798	41.78	<.0001

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	-9.432039841 B	1.58678711	-5.94	<.0001
POND 1	0.054134365 B	0.15467248	0.35	0.7273
POND 2	0.007033449 B	0.11165526	0.06	0.9499
POND 3	-0.174302587 B	0.12868697	-1.35	0.1796
POND 4	-0.018521732 B	0.11235126	-0.16	0.8695
POND 5	-0.126419116 B	0.10805039	-1.17	0.2457
POND 6	-0.026660389 B	0.10516802	-0.25	0.8006
POND 7	-0.033509202 B	0.09099004	-0.37	0.7137
POND 8	-0.086771999 B	0.11144842	-0.78	0.4386
POND 9	-0.033014208 B	0.12162730	-0.27	0.7868
POND 10	0.000000000 B	.	.	.
LSVL	5.134435845	0.79439065	6.46	<.0001

NOTE: The X'X matrix has been found to be singular, and a generalized inverse was used to solve the normal equations. Terms whose estimates are followed by the letter 'B' are not uniquely estimable.

Chapter 6

Determinants of the Grade Point Average

This exercise uses the *grades* dataset.

SAS COMMANDS FOR BOX 15.8 Analysis of the Grades dataset

```
Commands  proc glm data=gandh.Chapter6;
           class YEAR;
           model GPA = YEAR VERBAL MATH / solution;
           run;
```

Menu route Statistics > Anova > Linear Models...
 GPA → Dependent
 YEAR → Class
 VERBAL MATH → Quantitative

Statistics

Parameter estimates

SAS OUTPUT FOR BOX 15.8 Analysis of the Grades dataset

The GLM Procedure

Dependent Variable: GPA

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	3	8.62390974	2.87463658	9.65	<.0001
Error	196	58.39609026	0.29793924		
Corrected Total	199	67.02000000			

R-Square	Coeff Var	Root MSE	GPA Mean
0.128677	20.75430	0.545838	2.630000

Source	DF	Type I SS	Mean Square	F Value	Pr > F
YEAR	1	1.15520000	1.15520000	3.88	0.0504
VERBAL	1	6.75953191	6.75953191	22.69	<.0001
MATH	1	0.70917782	0.70917782	2.38	0.1245

Source	DF	Type III SS	Mean Square	F Value	Pr > F
YEAR	1	0.84599265	0.84599265	2.84	0.0936
VERBAL	1	5.15998884	5.15998884	17.32	<.0001
MATH	1	0.70917782	0.70917782	2.38	0.1245

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	0.5930373541 B	0.43943013	1.35	0.1787
YEAR 1	0.1304169627 B	0.07739531	1.69	0.0936
YEAR 2	0.0000000000 B	.	.	.
VERBAL	0.0022879692	0.00054978	4.16	<.0001
MATH	0.0009374861	0.00060765	1.54	0.1245

NOTE: The X'X matrix has been found to be singular, and a generalized inverse was used to solve the normal equations. Terms whose estimates are followed by the letter 'B' are not uniquely estimable.

Chapter 7

Weight, fat and sex

This question requires a factorial analysis, with an interaction between a categorical and continuous explanatory variable.

SAS COMMANDS FOR BOX 15.9 **Analysis for weight, fat and sex**

Commands `proc glm data=gandh.Chapter7;`
`class SEX;`
`model FAT = SEX|WEIGHT / solution;`
`run;`

Menu route `Statistics > Anova > Linear Models...`
`FAT` → Dependent
`SEX` → Class
`WEIGHT` → Quantitative
`Model`
`Independent` → Effects in Model
`SEX WEIGHT` → Add
`(SEX & WEIGHT)` → Cross
`Statistics`
 Parameter Estimates
`Tests`
 Type 1

SAS OUTPUT FOR BOX 15.9 **Analysis for weight, fat and sex**

```

                                The GLM Procedure

Dependent Variable: FAT

Source              DF          Sum of
                    Squares      Mean Square   F Value    Pr > F
Model                3          188.2834083      62.7611361   31.24    <.0001
Error               15          30.1376444       2.0091763
Corrected Total     18          218.4210526

                    R-Square      Coeff Var      Root MSE      FAT Mean
                    0.862020      4.996592      1.417454      28.36842

Source              DF          Type I SS      Mean Square   F Value    Pr > F
SEX                 1          90.32105263     90.32105263   44.95    <.0001
WEIGHT              1          87.10492012     87.10492012   43.35    <.0001
WEIGHT*SEX         1          10.85743552     10.85743552    5.40    0.0345

Source              DF          Type III SS     Mean Square   F Value    Pr > F
SEX                 1          2.10773879      2.10773879    1.05    0.3220
WEIGHT              1          79.54187398     79.54187398   39.59    <.0001
WEIGHT*SEX         1          10.85743552     10.85743552    5.40    0.0345

Parameter          Estimate          Standard
                    Error      t Value    Pr > |t|
Intercept          11.57095793 B    2.86807619     4.03    0.0011
SEX 1              -6.33128851 B    6.18148377    -1.02    0.3220
SEX 2               0.00000000 B    .              .        .
WEIGHT             0.18550431 B    0.03567800     5.20    0.0001
WEIGHT*SEX 1       0.21738825 B    0.09351504     2.32    0.0345
WEIGHT*SEX 2       0.00000000 B    .              .        .

```

NOTE: The X'X matrix has been found to be singular, and a generalized inverse was used to solve the normal equations. Terms whose estimates are followed by the letter 'B' are not uniquely estimable.

Chapter 8

Combining data from different experiments

We assume that you use Excel or some other software to stack the columns on top of each other into a new dataset. (A rather laborious way in SAS is to go into edit mode and edit the data table directly. Beware that there is another variable called YEAR in the chapter 8 dataset. If you find an easy way in SAS, do let us know!)

SAS COMMANDS FOR BOX 15.11 **Bird data combined**

```

Commands  proc glm data=gandh.Chapter8;
           class YEAR;
           model YOUNG = YEAR SONGDAY /solution;
           run;

```

Menu route Statistics > Anova > Linear Models...

YOUNG → Dependent

YEAR → Class

SONGDAY → Quantitative

Statistics

Type I

Parameter estimates

SAS OUTPUT FOR BOX 15.11 **Bird data combined**

The GLM Procedure

Dependent Variable: YOUNG

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	22.90003809	4.58000762	3.76	0.0073
Error	38	46.28178009	1.21794158		
Corrected Total	43	69.18181818			

R-Square	Coeff Var	Root MSE	YOUNG Mean
0.331012	38.53855	1.103604	2.863636

Source	DF	Type I SS	Mean Square	F Value	Pr > F
YR	4	10.69769120	2.67442280	2.20	0.0879
SONGDAY	1	12.20234690	12.20234690	10.02	0.0030

Source	DF	Type III SS	Mean Square	F Value	Pr > F
YR	4	19.78164888	4.94541222	4.06	0.0077
SONGDAY	1	12.20234690	12.20234690	10.02	0.0030

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	3.952056192 B	0.68156014	5.80	<.0001
YR 1	1.717708704 B	0.77197644	2.23	0.0321
YR 2	2.949080600 B	0.78840872	3.74	0.0006
YR 3	1.048516981 B	0.61473929	1.71	0.0962
YR 4	1.985456395 B	0.80929746	2.45	0.0189
YR 5	0.000000000 B	.	.	.
SONGDAY	-0.109127384	0.03447667	-3.17	0.0030

NOTE: The X'X matrix has been found to be singular, and a generalized inverse was used to solve the normal equations. Terms whose estimates are followed by the letter 'B' are not uniquely estimable.

Chapter 9

Stabilising the variance in a blocked experiment

The commands required for the first part of this exercise a) transform the data; b) perform the analysis on the transformed data, storing the residuals and plotting the residuals against the fitted values; c) display descriptive statistics for the residuals by treatment. We do all three transformations and analyses at once, illustrating SAS's flexibility in this respect.

However, there is one way in which SAS is not flexible. The chapter nine worksheet contains datasets of many different lengths. If you use the command line route for this exercise, SAS will not provide you with summary statistics for the residuals of treatment five—the mixed lengths of variables causing SAS some confusion. Therefore, for this exercise, we have imported the *cotyledons* dataset separately.

SAS COMMANDS FOR BOXES 15.12 TO 15.17, AND FIGURE 15.6 TO 15.8

```
Commands  data;
           set gandh.cotyledons;
           SQRTNCOT =sqrt(NCOT);
           LOGNCOT=log(NCOT);
           INVNCOT=1/ NCOT;

           proc glm;
             class BLOCK TRMNT;
             model SQRTNCOT LOGNCOT INVNCOT = BLOCK TRMNT;
             output p= SQRTfv LOGfv INVfv STUDENT= SQRTres LOGres
                   INVres;

           proc gplot;
             plot SQRTres * SQRTfv;
             plot LOGres * LOGfv;
             plot INVres * INVfv;

           proc means;
             by TRMNT;
             var SQRTres LOGres INVres;
           run;
```

(Contd.)

SAS COMMANDS FOR BOXES 15.12 TO 15.17, AND FIGURE 15.6 TO 15.8 (Contd.)

Menu route Edit > Mode > Edit...

Data > Transform > Compute...

Type "SQRTNCOT" into top left box

sqrt(NCOT) → Main pane

Data > Transform > Compute...

Type "LOGNCOT" into top left box

log(NCOT) → Main pane

Data > Transform > Compute...

Type "INVNCOT" into top left box

1/NCOT → Main pane

Statistics > Anova > Linear Models...

SQRTNCOT LOGNCOT INVNCOT → Dependent

BLOCK TRMNT → Class

Plots with Residual tab

Plot residuals vs variables

Residuals: Studentized

Variables: Predicted Y

Save Data

Create and save diagnostic data

STUDENT → Add

Now click on the most recent 'Diagnostic Table' icon in left hand pane of Analyst, and choose File > Save as by SAS Name..., saving it under some name. Then open that file within Analyst using File > Open by SAS Name.

Statistics > Descriptive > Summary Statistics...

_STUDENT1 _STUDENT2 _STUDENT3 → Analysis

TRMNT → Class

Just the first ANOVA table and residual plots is illustrated below:

SAS OUTPUT FOR BOX 15.12 **Analysis of blocked experiment**

The GLM Procedure

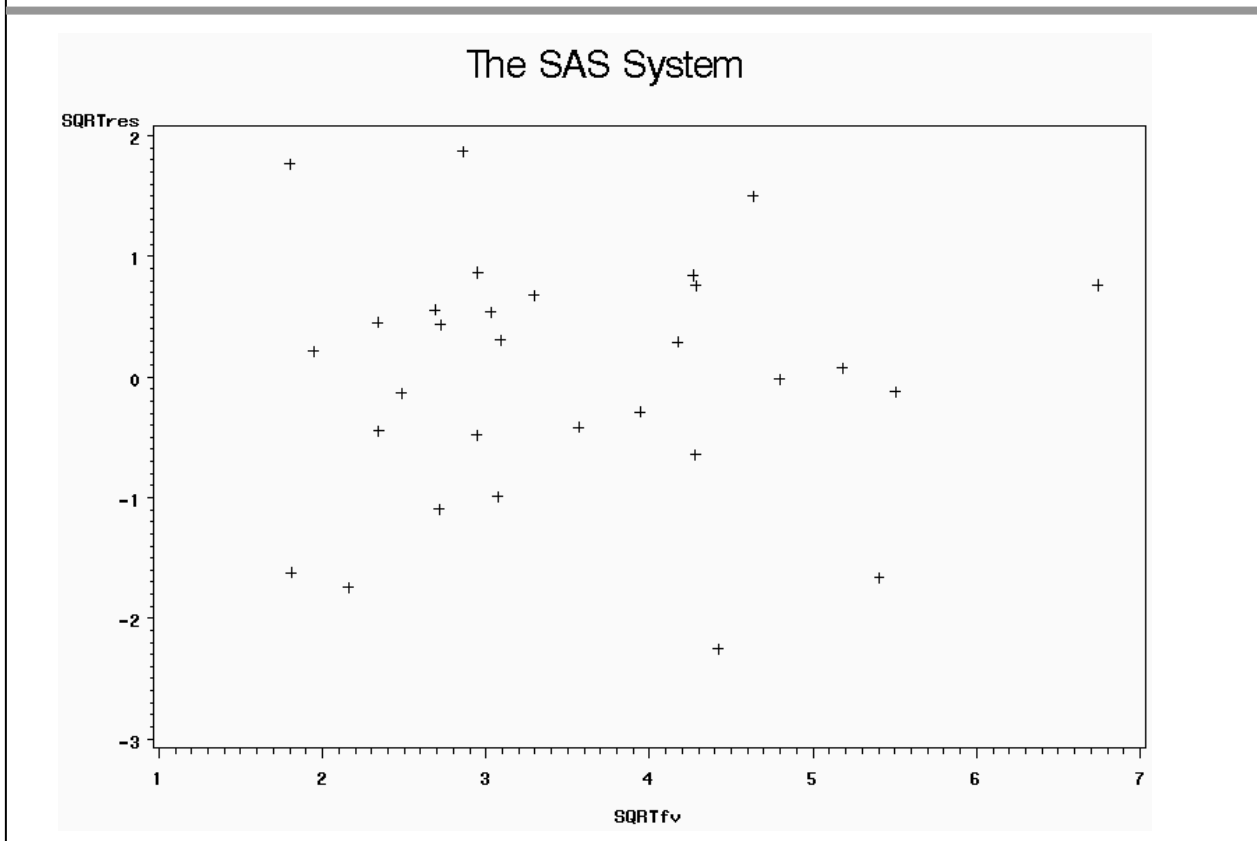
Dependent Variable: SQRTNCOT

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	9	44.44384413	4.93820490	55.19	<.0001
Error	20	1.78947158	0.08947358		
Corrected Total	29	46.23331571			

R-Square	Coeff Var	Root MSE	SQRTNCOT Mean
0.961295	8.508592	0.299121	3.515521

Source	DF	Type I SS	Mean Square	F Value	Pr > F
BLOCK	5	23.90980329	4.78196066	53.45	<.0001
TRMNT	4	20.53404083	5.13351021	57.37	<.0001

Source	DF	Type III SS	Mean Square	F Value	Pr > F
BLOCK	5	23.90980329	4.78196066	53.45	<.0001
TRMNT	4	20.53404083	5.13351021	57.37	<.0001

SAS OUTPUT FOR FIGURE 15.6 **Standardised residual plot**

SAS OUTPUT FOR BOX 15.13, 15.15 AND 15.17 Examining the residuals					
The SAS System					
----- TRMNT=1 -----					
The MEANS Procedure					
Variable	N	Mean	Std Dev	Minimum	Maximum
SQRTres	6	3.145632E-16	0.9364663	-1.6633021	0.8675659
LOGres	6	1.082467E-15	0.7729623	-1.1574905	0.9904145
INVres	6	3.666049E-16	0.6576565	-0.7034478	1.0938756
----- TRMNT=2 -----					
Variable	N	Mean	Std Dev	Minimum	Maximum
SQRTres	6	-3.33067E-16	1.0912566	-1.0959101	1.8692103
LOGres	6	8.881784E-16	0.8702716	-1.1118747	1.3163716
INVres	6	9.251859E-18	0.6243825	-0.7704611	0.9619229
----- TRMNT=3 -----					
Variable	N	Mean	Std Dev	Minimum	Maximum
SQRTres	6	-1.21662E-15	0.4247971	-0.4433083	0.5573088
LOGres	6	7.031412E-16	0.4116188	-0.4166356	0.6015208
INVres	6	-1.75785E-16	0.3131422	-0.3915120	0.5031805
----- TRMNT=4 -----					
Variable	N	Mean	Std Dev	Minimum	Maximum
SQRTres	6	-4.57967E-16	1.4024633	-1.7435106	1.7666929
LOGres	6	-6.4763E-17	1.8411957	-2.5521149	1.9012582
INVres	6	-7.17019E-17	2.0836062	-3.4601347	1.9890440
----- TRMNT=5 -----					
Variable	N	Mean	Std Dev	Minimum	Maximum
SQRTres	6	-5.9952E-15	1.3359768	-2.2548799	1.4984861
LOGres	6	7.21645E-16	1.0419813	-1.0269795	1.3153124
INVres	6	-1.06396E-15	0.8591633	-1.0726851	1.2975456

The commands also produce the output for the log and inverse transformations.

Checking the 'perfect' model

The commands to produce the histograms of the original and transformed data are reproduced below. We specify 'midpoints' so that SAS will use the same scale on the *x*-axes, making the histograms easier to compare by eye.

SAS COMMANDS TO PRODUCE FIGURE 15.9 A AND B

```
Commands  data;
           set gandh.chapter9;
           proc capability noprint;
             var MALE FEMALE;
             histogram / midpoints = 0.2 to 1.1 by 0.1;
           run;
```

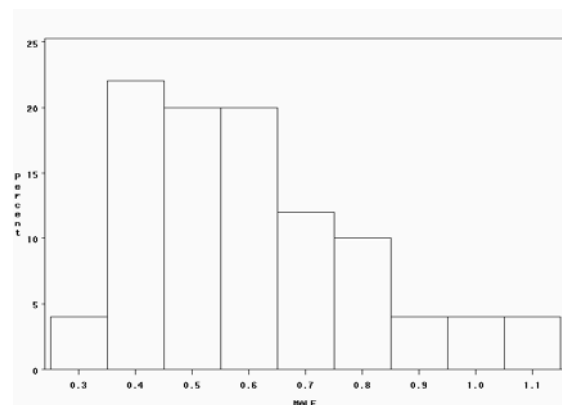
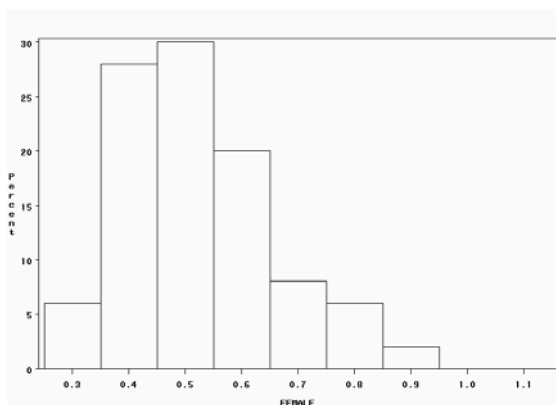
Menu route Graphs > Histogram...

MALE FEMALE → Analysis

Display

Type 0.2 , 1.1, 0.1 into: Midpoints of histogram intervals

SAS OUTPUT FOR FIGURE 15.9 A AND B **Histograms of squirrel body weights**



SAS COMMANDS TO PRODUCE FIGURE 15.9 C AND D

```

Commands  data ;
           set gandh.Chapter9;
           LMALE =log(MALE);
           LFEM =log(FEMALE);
           run;

           proc capability noprint;
             var LMALE LFEM;
             histogram / midpoints = -1.3 to 0.1 by 0.1;
           run;

```

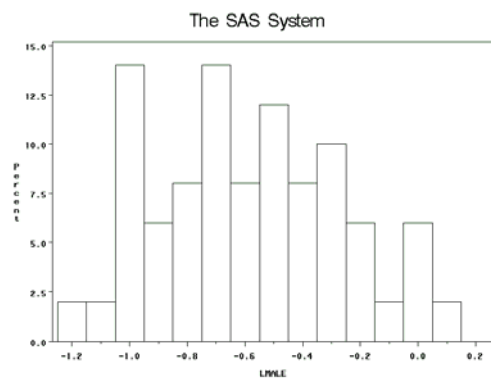
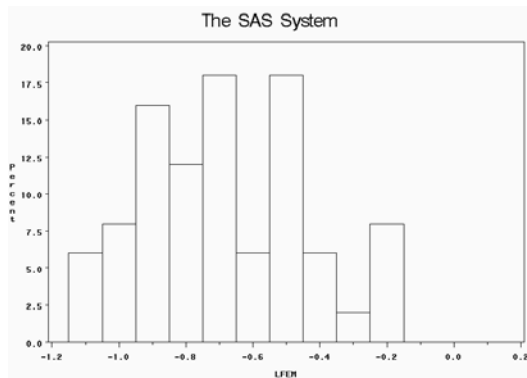
Menu route Edit > Mode > Edit
 Data > Transform > Compute...
 Type "LMALE" into top left box
 log(MALE) → Main pane

Data > Transform > Compute...
 Type "LFEM" into top left box
 log(FEMALE) → Main pane

Graphs > Histogram...
 LMALE LFEM → Analysis

Display

Type -1.3, 0.1, 0.1 into: Midpoints of histogram intervals

SAS OUTPUT FOR FIGURE 15.9 C AND D **Histograms of squirrel log (body weights)**

We now produce the Normal probability plots. We use the option 'normaltest' so that SAS will produce tests for Normality in among all the other printed output. These include the Anderson-Darling test shown in the main text.

SAS COMMANDS FOR FIGURE 15.10

```

Commands  data;
           set gandh.chapter9;
           LMALE =log(MALE);
           LFEM  =log(FEMALE);

           proc capability normaltest;
           var MALE FEMALE LMALE LFEM;
           probplot / normal(mu=est sigma=est);
           run;

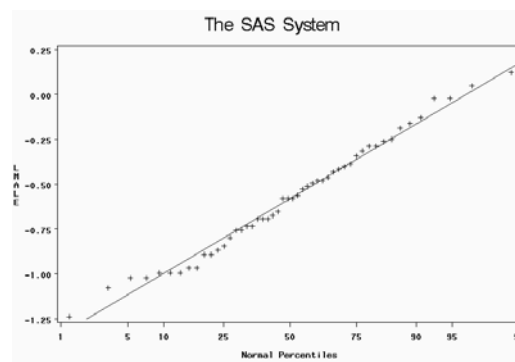
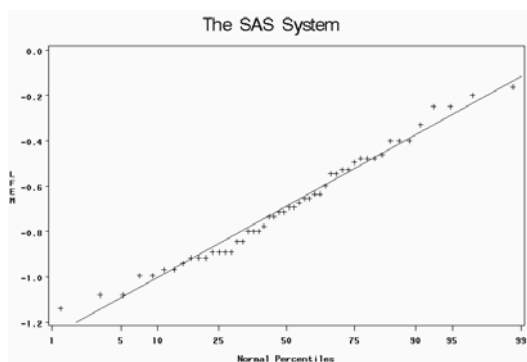
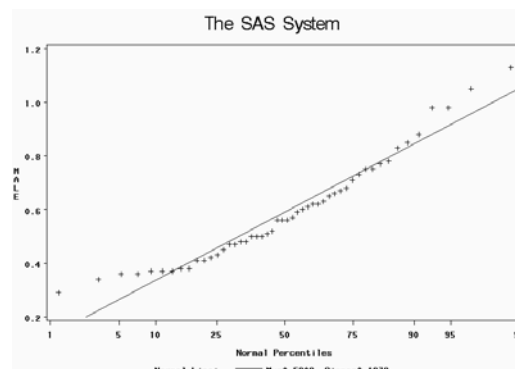
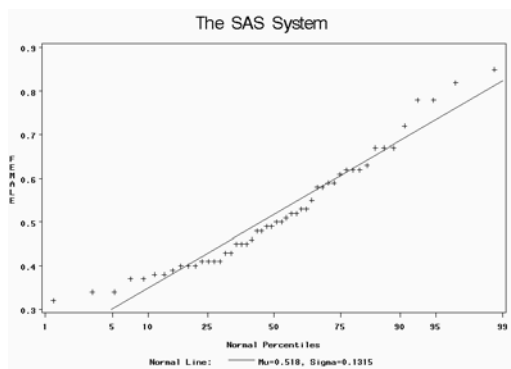
```

Menu route Graphs > Probability Plot...

MALE FEMALE LMALE LFEM → Analysis

This produces the graphs below.

SAS OUTPUT FOR FIGURE 15.10 Normal probability plots for the squirrels dataset



When comparing these graphs with those in the main text, note that the axes are reversed. The Anderson-Darling Normality test provides us with a criterion for rejecting the assumption of Normality. The p -value is below 0.05 for the raw data, but not for the transformed data.

EDITED SAS OUTPUT SHOWING THE ANDERSON-DARLING TEST FOR NORMALITY

```

The CAPABILITY Procedure
  Variable:  MALE
  Tests for Normality

Test                --Statistic---    -----p-value-----
Shapiro-Wilk        W      0.940648      Pr < W      0.014
Kolmogorov-Smirnov  D      0.099719      Pr > D      >0.150
Cramer-von Mises    W-Sq   0.115478      Pr > W-Sq   0.071
Anderson-Darling    A-Sq   0.812115      Pr > A-Sq   0.035

```

```

The CAPABILITY Procedure
  Variable:  FEMALE
  Tests for Normality

Test                --Statistic---    -----p-value-----
Shapiro-Wilk        W      0.941389      Pr < W      0.015
Kolmogorov-Smirnov  D      0.103651      Pr > D      >0.150
Cramer-von Mises    W-Sq   0.128795      Pr > W-Sq   0.045
Anderson-Darling    A-Sq   0.858657      Pr > A-Sq   0.025

```

```

The CAPABILITY Procedure
  Variable:  LMALE
  Tests for Normality

Test                --Statistic---    -----p-value-----
Shapiro-Wilk        W      0.982436      Pr < W      0.658
Kolmogorov-Smirnov  D      0.064466      Pr > D      >0.150
Cramer-von Mises    W-Sq   0.029189      Pr > W-Sq   >0.250
Anderson-Darling    A-Sq   0.236910      Pr > A-Sq   >0.250

```

```

The CAPABILITY Procedure
  Variable:  LFEM
  Tests for Normality

Test                --Statistic---    -----p-value-----
Shapiro-Wilk        W      0.974443      Pr < W      0.347
Kolmogorov-Smirnov  D      0.096994      Pr > D      >0.150
Cramer-von Mises    W-Sq   0.055795      Pr > W-Sq   >0.250
Anderson-Darling    A-Sq   0.370725      Pr > A-Sq   >0.250

```

The following commands allow you to create the null datasets designed to follow a Normal distribution, with a mean and standard deviation equal to that of the female squirrel weights (which can be obtained using PROC MEANS, as used earlier).

SAS COMMANDS TO SIMULATE DATASETS	
Commands	<pre> data; do I=1 to 50; noise=normal(0); SIMFEM1=0.518+0.1315*noise; output; end; run; </pre>
Menu route	<p>File > New (<i>necessary only the first time</i>)</p> <p>Data > Random Variates > Normal</p> <p>50 → Number of values to generate (<i>necessary only the first time</i>)</p> <p>SIMFEM1 → New column name (<i>change the name each time</i>)</p> <p>0.518 → mean</p> <p>0.1315 → standard deviation</p>

This can be repeated to create as many columns of simulated female (or male) data as your require. This is particularly easy using the written commands in the session window, as each time you press the submit button, a new dataset is produced in the work folder (called data1, data2 etc.). These datasets may then be examined using the histogram and Normal probability plot commands used earlier.

Chapter 10

Partitioning a sum of squares into polynomial components

Note that in the command language the analysis is conducted in two parts. In the first part, the 'interaction means' are stored at the same time as conducting the PROC GLM and printing out a table of means. In the second part, we set up some formats using PROC GOPTIONS, and then plot using PROC GPLOT. There are three SYMBOLi statements because there are three lines.

Notice that in the menu route we first remove BSPACE and BVARIETY from the model. This is necessary in order to control the order of variables in the model. The menu route allows us to put error bars on the plot, whereas the command route does not.

SAS COMMANDS FOR BOX 15.18 AND FIGURE 10.8

```

Commands  proc glm data=gandh.Chapter10;
           class BBLOCK BSPACE BVARIETY;
           model BYIELD = BBLOCK BSPACE|BVARIETY;
           lsmeans BSPACE*BVARIETY / out=WORK.INDIAG2;

           goptions reset=SYMBOL;

           proc goptions;
             SYMBOL1 LINE=1 CI=BLACK INTERPOL=JOIN;
             SYMBOL2 LINE=2 CI=BLUE INTERPOL=JOIN;
             SYMBOL3 LINE=3 CI=RED INTERPOL=JOIN;

           proc gplot data=WORK.INDIAG2;
             label lsmean = "Predicted YIELD";
             plot lsmean * BSPACE = BVARIETY;
           run;

```

Menu route Statistics > Anova > Linear Models...

BYIELD → Dependent

BBLOCK BSPACE BVARIETY → Class

Model

Effects in Model

BSPACE BVARIETY → Remove

Select 2 as factorial order by clicking on ▲

Independent → Effects in model

(BSPACE & BVARIETY) → Factorial

Plots

with Means tab

Plot dependent means for two-way effects

Predicted means

2 se

Means

with LSMeans tab

BSPACE*BVARIETY → LS Mean

SAS OUTPUT FOR BOX 15.18 **Analysis for barley yield**

The GLM Procedure

Dependent Variable: BYIELD

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	11	2203.527778	200.320707	11.34	<.0001
Error	24	424.111111	17.671296		
Corrected Total	35	2627.638889			

R-Square	Coeff Var	Root MSE	BYIELD Mean
0.838596	7.272180	4.203724	57.80556

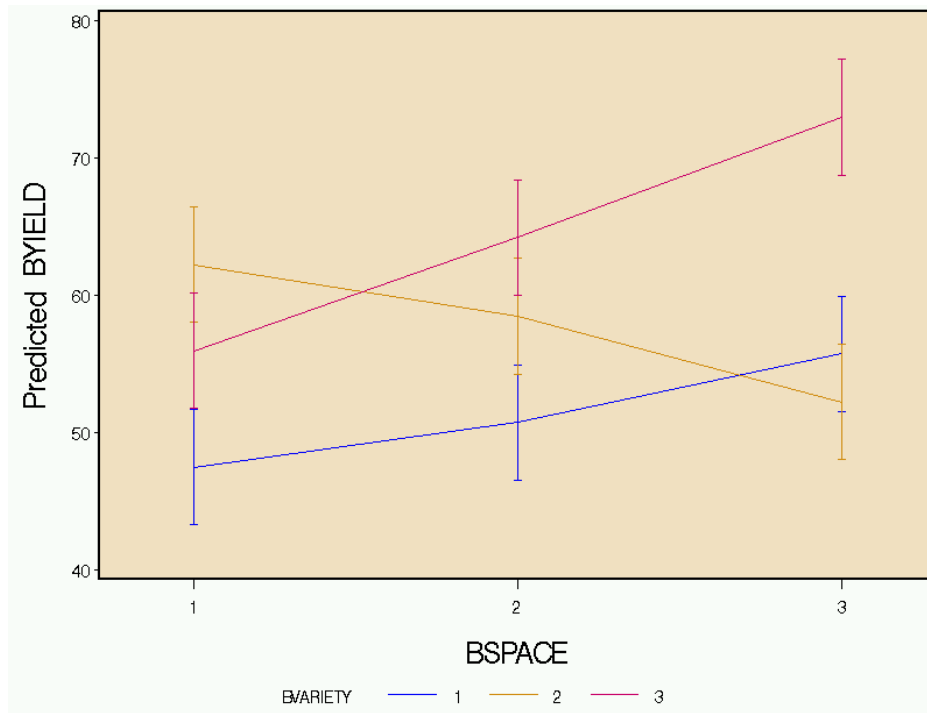
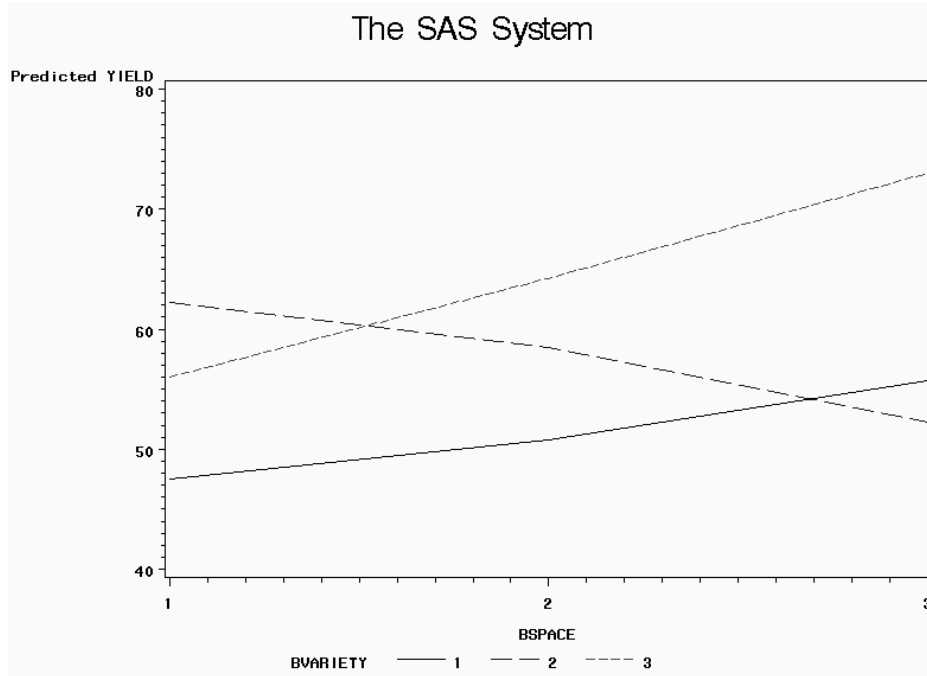
Source	DF	Type III SS	Mean Square	F Value	Pr > F
BBLOCK	3	255.638889	85.212963	4.82	0.0091
BSPACE	2	155.055556	77.527778	4.39	0.0238
BVARIETY	2	1027.388889	513.694444	29.07	<.0001
BSPACE*BVARIETY	4	765.444444	191.361111	10.83	<.0001

The GLM Procedure
Least Squares Means

BSPACE	BVARIETY	BYIELD LSMEAN
1	1	47.5000000
1	2	62.2500000
1	3	56.0000000
2	1	50.7500000
2	2	58.5000000
2	3	64.2500000
3	1	55.7500000
3	2	52.2500000
3	3	73.0000000

SAS OUTPUT FOR FIGURE 10.8

Interaction diagram from the *barley* dataset. Above by the command line route, and below by the menu route



A second analysis is then conducted on the *barley* dataset, involving a polynomial decomposition. We choose to include the interaction of BVARIETY with the quadratic term BSPACE*BSPACE so that the polynomial analysis is an exact decomposition of the non-polynomial analysis. This means that the error SS, MS and DF will be the same in the two analysis, allowing us to compare them more easily. It also allows us to investigate whether the response of yield to spacing is curvilinear for some varieties but not others. Note the use of Type 1 (that is, sequential) sums of squares, as the use of BSPACE as a covariate has destroyed some of the orthogonality of the design. For example, knowing BSPACE certainly helps in guessing the value of BSPACE*BSPACE!

SAS COMMANDS FOR BOX 15.19 **Second analysis for barley yield**

```
Commands  proc glm data=gandh.Chapter10;
           class BBLOCK BSPACE BVARIETY;
           model BYIELD = BBLOCK BVARIETY|BSPACE|BSPACE / SS1;
           run;
```

Menu route Statistics > Anova > Linear Models...

BYIELD → Dependent

BBLOCK BVARIETY → Class

BSPACE → Quantitative

Model

Effects in Model

BSPACE BVARIETY → Remove

Select 2 as Polynomial order by clicking on ▲

Independent → Effects in model

BVARIETY → Add

BSPACE → Add

(BSPACE & BVARIETY) → Cross

BSPACE → Polynomial

(BSPACE & BVARIETY) → Cross

Select BVARIETY in 'Independent' pane, and BSPACE*

BSPACE in 'Effects in model' pane, then click Cross

Statistics

Type 3

Type 1

SAS OUTPUT FOR BOX 15.19 Second analysis for barley yield					
The GLM Procedure					
Dependent Variable: BYIELD					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	11	2203.527778	200.320707	11.34	<.0001
Error	24	424.111111	17.671296		
Corrected Total	35	2627.638889			
	R-Square	Coeff Var	Root MSE	BYIELD Mean	
	0.838596	7.272180	4.203724	57.80556	
Source	DF	Type I SS	Mean Square	F Value	Pr > F
BBLOCK	3	255.638889	85.212963	4.82	0.0091
BVARIETY	2	1027.388889	513.694444	29.07	<.0001
BSPACE	1	155.041667	155.041667	8.77	0.0068
BSPACE*BVARIETY	2	759.083333	379.541667	21.48	<.0001
BSPACE*BSPACE	1	0.013889	0.013889	0.00	0.9779
BSPACE*BSPACE*BVARIE	2	6.361111	3.180556	0.18	0.8364

Chapter 11

Finding the best treatment for cat fleas

The final model is:

SAS COMMANDS FOR BOX 15.20 Analysis for cat fleas	
Commands	<pre>proc glm data=gandh.Chapter11; class TRTMT CARPET; model LOGFLEAS = TRTMT NCATS CARPET / SS3; run;</pre>
Menu route	Statistics > Anova > Linear Models... LOGFLEAS → Dependent TRTMT CARPET → Class NCATS → Quantitative

SAS OUTPUT FOR BOX 15.20 **Analysis for cat fleas**

The GLM Procedure						
Dependent Variable: LOGFLEAS						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	3	31.44462221	10.48154074	16.57	<.0001	
Error	85	53.75863885	0.63245457			
Corrected Total	88	85.20326106				
	R-Square	Coeff Var	Root MSE	LOGFLEAS Mean		
	0.369054	19.76603	0.795270	4.023419		
Source	DF	Type III SS	Mean Square	F Value	Pr > F	
TRTMT	1	5.75019579	5.75019579	9.09	0.0034	
NCATS	1	22.96168694	22.96168694	36.31	<.0001	
CARPET	1	6.10297060	6.10297060	9.65	0.0026	

Multiplicity of p-values

To create the random data set you can either write a short program in the command line route, or make use of the Random Variates option in menus.

SAS COMMANDS TO PRODUCE THE NULL DATASET

```

Commands  data WORK.randata;
          array COL{11};
          drop I J;
          do I=1 to 30;
            do J=1 to 11;
              COL{J}=10.5+2.4*normal(0);
            end;
          end;
          output;
          end;
          run;

```

Menu route File > New
 Data > Random Variates > Normal
 30 → Number of values to generate
 COL1 → New column name
 10.5 → mean
 2.4 → standard deviation
 Edit > Mode > Edit
 Data > Random Variates > Normal
 COL2 → New column name
 10.5 → mean
 2.4 → standard deviation

and repeat this final step for COL3, COL4 etc up to COL11

PROC REG will provide output giving the p -value for the regression as a whole, with the p -values for each individual X variable in the analysis being found in the table of coefficients. (The p -value resulting from the t test for the slope against zero for a particular variable is the same as the p -value for the F ratio test for the same explanatory variable in the analysis of variance table. This arises from the fact that $t^2 = F$; see Appendix 3).

SAS COMMANDS TO PRODUCE OUTPUT FOR TABLE 15.6

```
Commands  proc reg data=gandh.Chapter11;
           model COL1= COL2-COL11;
           run;
```

Menu route Statistics > Regression > Linear ...

COL1 → Dependent

COL2 COL3 COL4 COL5 COL6 COL7 COL8 COL9 COL10

COL11 → Explanatory

EXAMPLE SAS OUTPUT FOR TABLE 15.6 Multiple regression with unrelated variables

The REG Procedure
Model: MODEL1
Dependent Variable: COL1

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	10	68.31104	6.83110	1.36	0.2349
Error	39	195.95819	5.02457		
Corrected Total	49	264.26922			

Root MSE	2.24156	R-Square	0.2585
Dependent Mean	10.23035	Adj R-Sq	0.0684
Coeff Var	21.91082		

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	1	6.90922	4.71496	1.47	0.1508
COL2	1	0.18923	0.14159	1.34	0.1891
COL3	1	-0.00625	0.14082	-0.04	0.9648
COL4	1	-0.26071	0.14593	-1.79	0.0818
COL5	1	0.07864	0.12610	0.62	0.5365
COL6	1	-0.13006	0.14481	-0.90	0.3746
COL7	1	0.30202	0.16153	1.87	0.0690
COL8	1	-0.06789	0.14262	-0.48	0.6367
COL9	1	0.25457	0.14110	1.80	0.0789
COL10	1	0.01550	0.14639	0.11	0.9162
COL11	1	-0.08342	0.14385	-0.58	0.5653

Chapter 12

How a nested analysis can solve problems of non-independence

These commands first calculate the variable LUPRATE, and then conduct a nested analysis.

SAS COMMANDS FOR BOX 15.21: *Nested analysis of sheep dataset*

```
Commands  data;
           set gandh.Chapter12;
           LUPRATE=NLOOKUPS/DURATION;

           proc mixed method=type1 covtest;
             class SEX SHEEP;
             model LUPRATE = SEX / htype=1 ddfm=satterth;
             random SHEEP(SEX);
           run;
```

Menu route Edit > Mode > Edit

Data > Transform > Compute...

Type "LUPRATE" into top left box

NLOOKUPS / DURATION → Main pane

Statistics > Anova > Mixed Models...

LUPRATE → Dependent

SEX SHEEP → Class

Model

SEX → Add

Random Effects

SHEEP → Add

Select SHEEP in Random effects box, and SEX in Class

box, and then click 'Nest'

Tests

Type III

Type I

Tests of variance components

Options

Type 1

SAS OUTPUT FOR BOX 15.21 **Nested analysis of sheep dataset**

The Mixed Procedure

Model Information

Data Set	WORK.DATA1
Dependent Variable	LUPRATE
Covariance Structure	Variance Components
Estimation Method	Type 1
Residual Variance Method	Factor
Fixed Effects SE Method	Model-Based
Degrees of Freedom Method	Satterthwaite

Class Level Information

Class	Levels	Values
SEX	2	1 2
SHEEP	3	1 2 3

Dimensions

Covariance Parameters	2
Columns in X	3
Columns in Z	6
Subjects	1
Max Obs Per Subject	120
Observations Used	120
Observations Not Used	0
Total Observations	120

Type 1 Analysis of Variance

Source	DF	Sum of Squares	Mean Square	Expected Mean Square
SEX	1	0.132816	0.132816	Var(Residual) + 20 Var(SHEEP(SEX)) + Q(SEX)

Type 1 Analysis of Variance

Source	Error Term	Error DF	F Value	Pr > F
SEX	MS(SHEEP(SEX))	4	1.80	0.2506

The Mixed Procedure

Type 1 Analysis of Variance

Source	DF	Sum of Squares	Mean Square	Expected Mean Square
SHEEP(SEX)	4	0.294774	0.073693	Var(Residual) + 20 Var(SHEEP(SEX))
Residual	114	0.312387	0.002740	Var(Residual)

Type 1 Analysis of Variance

Source	Error Term	Error DF	F Value	Pr > F
SHEEP(SEX)	MS(Residual)	114	26.89	<.0001
Residual

Covariance Parameter Estimates

Cov Parm	Estimate	Standard Error	Z Value	Pr Z
SHEEP(SEX)	0.003548	0.002606	1.36	0.1733
Residual	0.002740	0.000363	7.55	<.0001

Fit Statistics

-2 Res Log Likelihood	-339.9
AIC (smaller is better)	-335.9
AICC (smaller is better)	-335.8
BIC (smaller is better)	-336.4

Type 1 Tests of Fixed Effects

Effect	Num DF	Den DF	F Value	Pr > F
SEX	1	4	1.80	0.2506

Chapter 13

Fig trees in Costa Rica

Descriptive statistics are obtained using proc means, or via summary statistics as follows:

SAS COMMANDS FOR TABLE 15.7 Means and variances for fig tree sites

```

Commands  proc means data=gandh.Chapter13 mean std median q1 q3
           min max;
           class SITE;
           var NINDIVS;

           proc means data=gandh.Chapter13 mean std median q1 q3
           min max;
           var NINDIVS;
           run;

```

Menu route Statistics > Descriptive > Summary Statistics...
 NINDIVS → Analysis
 SITE → Class

Statistics > Descriptive > Summary Statistics...
 NINDIVS → Analysis

SAS OUTPUT FOR TABLE 15.7 Means and variances for fig tree sites

```

                                The MEANS Procedure
                                Analysis Variable : NINDIVS

  SITE      N      N      Mean      Std Dev      Minimum      Maximum
  -----
    1      100    100     9.5100000    3.2176753     2.0000000    17.0000000
    2      100    100    18.9300000    4.4090243     8.0000000    28.0000000
    3      100    100    41.4300000    6.2607564    27.0000000    56.0000000
  -----

                                The MEANS Procedure
                                Analysis Variable : NINDIVS

      N      Mean      Std Dev      Minimum      Maximum
  -----
    300    23.2900000    14.2394443     2.0000000    56.0000000
  -----

```

The appropriate glm analysis is then as follows:

SAS COMMANDS FOR BOX 15.22	
Commands	<pre> data; set gandh.Chapter13; SQRTN =sqrt(NINDIVS); proc glm; class SITE; model SQRTN = SITE; run; </pre>
Menu route	<p>Data > Transform > Compute...</p> <p>Type "SQRTN" into top left box</p> <p>sqrt(NINDIVS) → Main pane</p> <p>Statistics > Anova > Linear Models...</p> <p>SQRTN → Dependent</p> <p>SITE → Class</p>

SAS OUTPUT FOR BOX 15.22					
The GLM Procedure					
Dependent Variable: SQRTN					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	582.6257220	291.3128610	1091.99	<.0001
Error	297	79.2310695	0.2667713		
Corrected Total	299	661.8567915			
	R-Square	Coeff Var	Root MSE	SQRTN Mean	
	0.880290	11.24851	0.516499	4.591711	
Source	DF	Type I SS	Mean Square	F Value	Pr > F
SITE	2	582.6257220	291.3128610	1091.99	<.0001
Source	DF	Type III SS	Mean Square	F Value	Pr > F
SITE	2	582.6257220	291.3128610	1091.99	<.0001